

IMPACASTORINA - PASSWORD

Assumptionless bounds for random trees

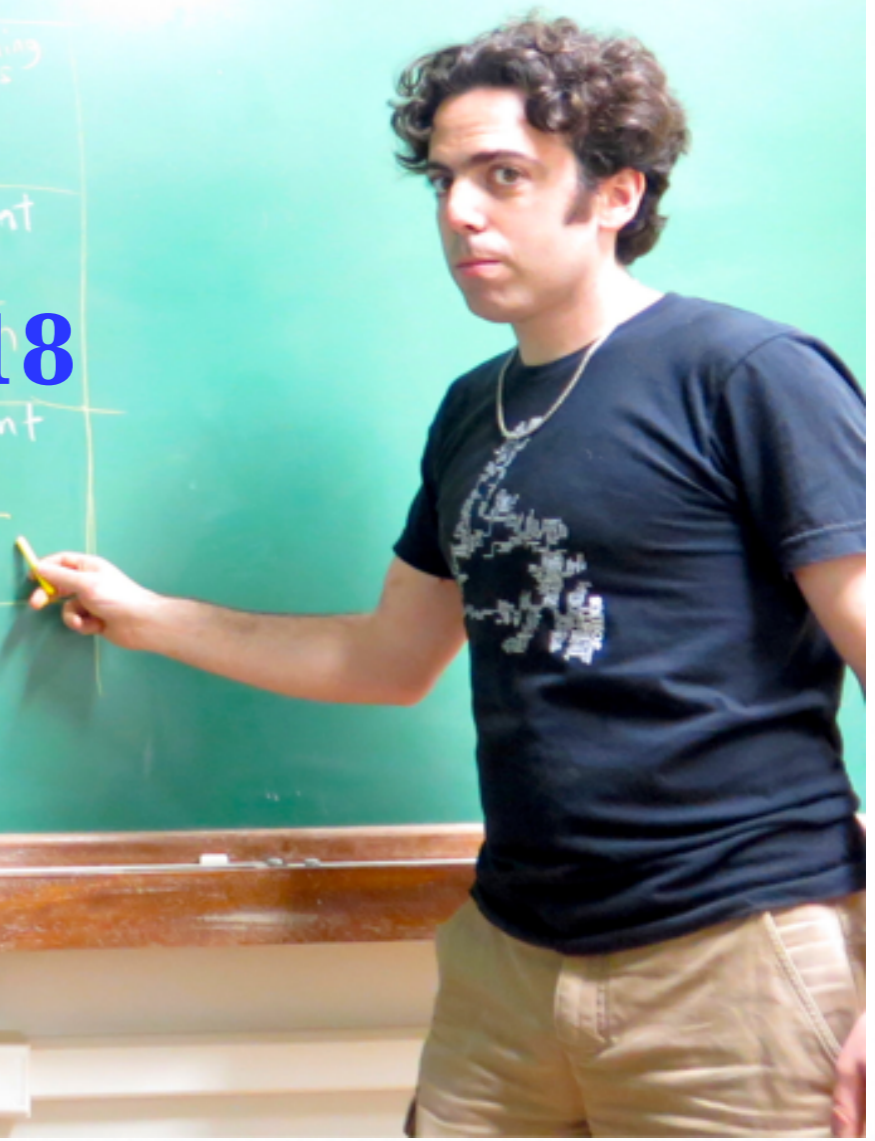
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Analysis of Algorithms 2018



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UNIVERSITET



Subcritical



Supercritical



Necessary
Sufficient

Symmetrization
+VE. Martingale
VC

Specific balanced
coloring + VC

$n \log^2 n$

Non-sudakov

n

$\log n$

Bid choice VC

Must kill initial

$\sim (n^c)$

RV + Make spanning
trees

$n \log n$

First moment

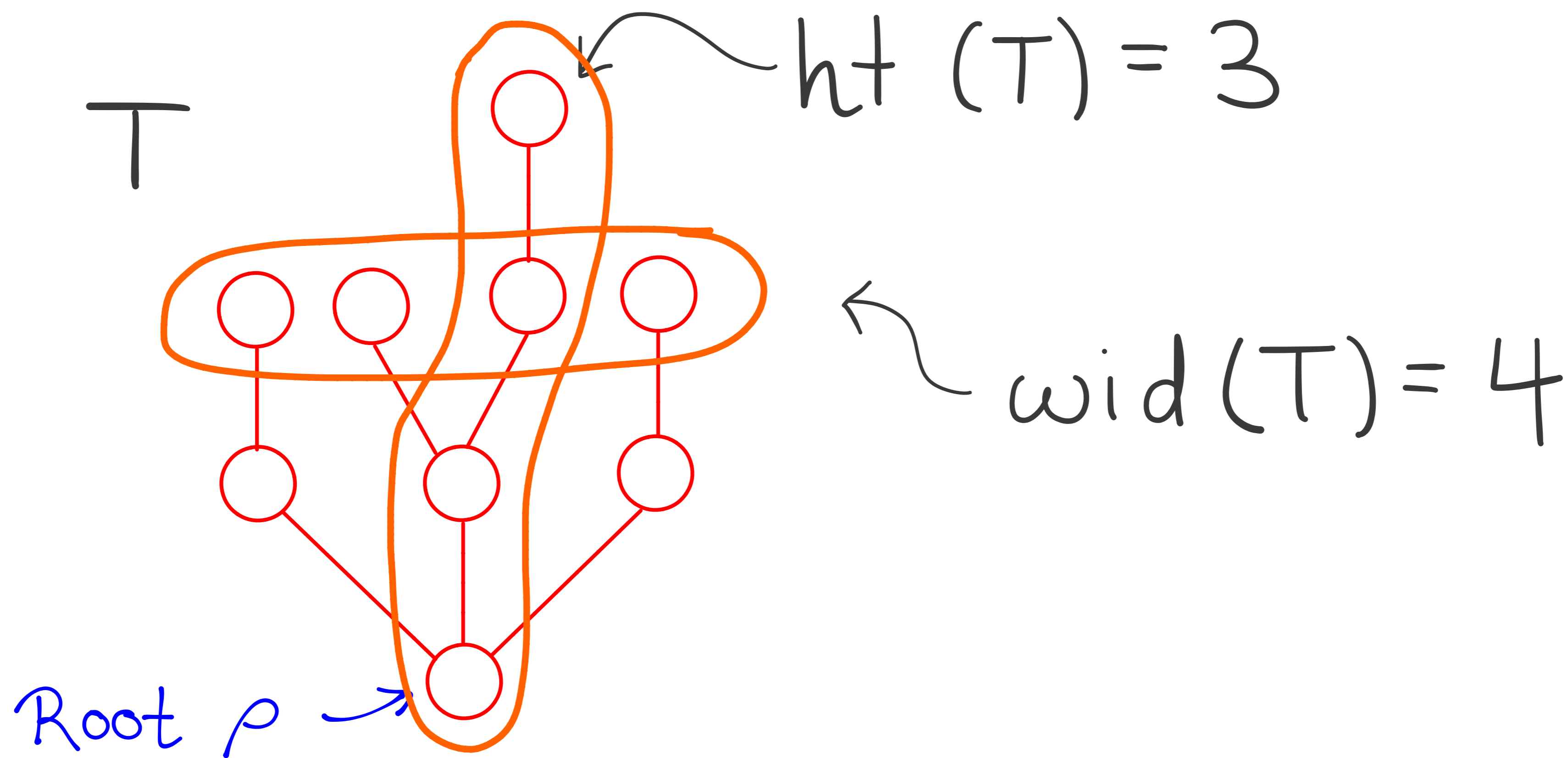
$(c-1) \log n$

Second moment

$(c-1) \log n$

$\log \log n$

Trees:

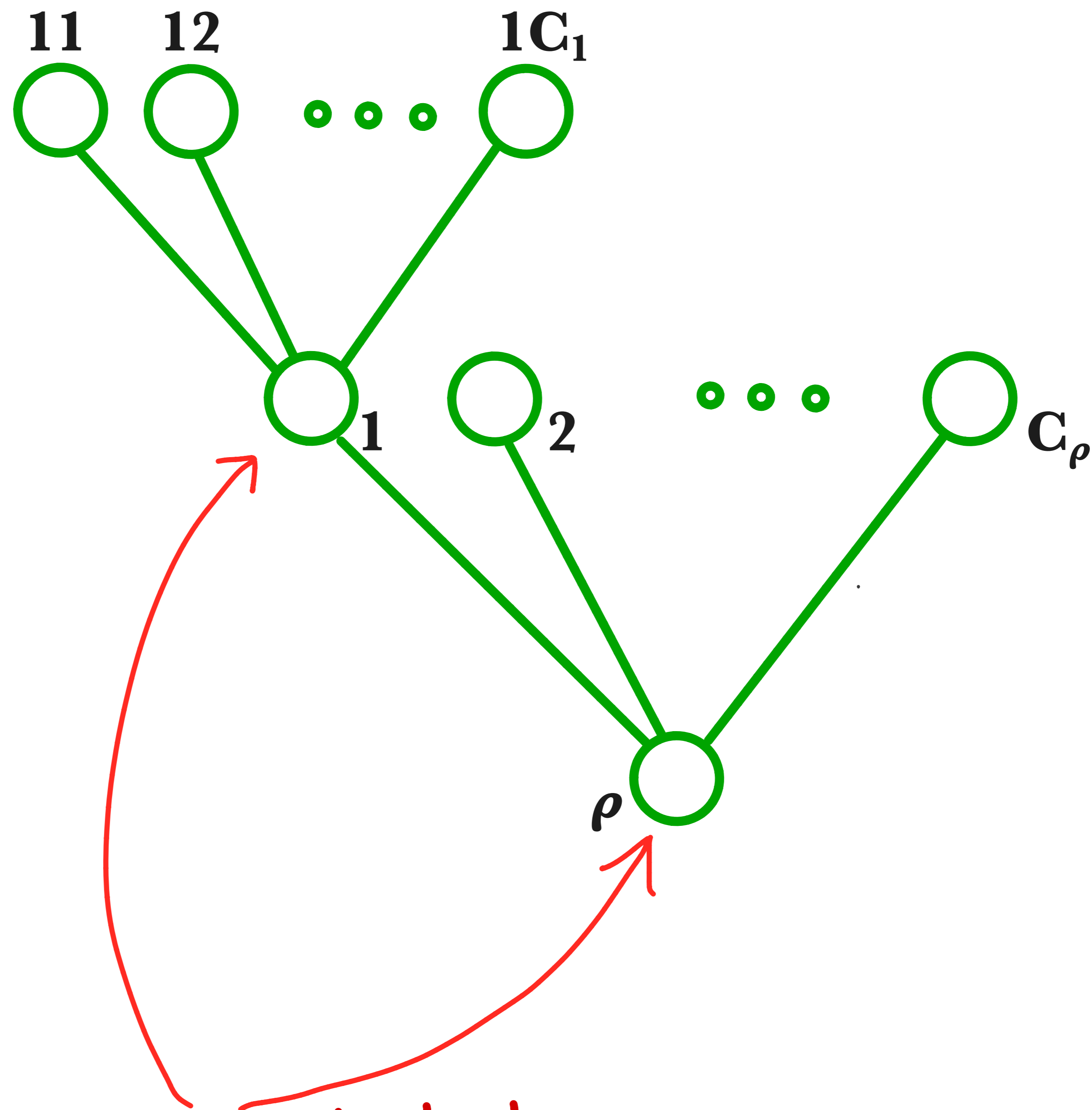


Height: Greatest distance from any node to the root } ht(T)

Width: Greatest # nodes on a single level. } wid(T)

Main Results Fix any prob. distribution $p = (p_0, p_1, p_2, \dots)$ on \mathbb{N} (so $\sum_{i \geq 0} p_i = 1$)

Let T be $GW(p)$ distributed.



children is distributed as p ,
independently at each node.

Main Results Fix any prob. distribution $p = (p_0, p_1, p_2, \dots)$ on \mathbb{N} (so $\sum_{i \geq 0} p_i = 1$)
Let T be $\text{GW}(p)$ distributed. Write $\mu(p) = \sum_{i \geq 0} i p_i \in [0, \infty]$.

Theorem ("Most trees are short & fat")

There is a universal constant $\delta > 0$ s.t.

$$\mathbb{P}(\text{ht}(T) > \frac{k}{1-p_1} \cdot \text{wid}(T)) \leq \exp(-\delta k).$$

Remark: Let $\sigma = \#$ nodes of T . If $\mu(p) > 1$ then $\mathbb{P}(\sigma = \infty) > 0$, and
 $\mathbb{P}(\text{ht}(T) = \text{wid}(T) = \infty \mid \sigma = \infty) = 1$.

Also, given that $\sigma < \infty$, the cond. dist. of T is $\text{GW}(\hat{p})$ where $\hat{p}_i = p_i$,

$$\mu(\hat{p}) \leq 1 \text{ so can assume } \mu(p) \leq 1.$$

Heuristic: GW trees satisfy $\text{wid}(T) \cdot \text{ht}(T) \approx \sigma$

$$\text{Implies "ht} > C^2 \cdot \text{wid" } \approx \text{"ht}^2 \geq C^2 \cdot \sigma \text{"}$$

$$\text{So might guess that } \mathbb{P}(\text{ht}(T) > \frac{k}{\sqrt{1-p_1}} \sqrt{\sigma}) \leq \exp(-\delta k^2)$$

Main Results Fix any prob. distribution $p = (p_0, p_1, p_2, \dots)$ on \mathbb{N} (so $\sum_{i \geq 0} p_i = 1$)
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Implies " $\text{ht} > C^2 \cdot \text{wid}$ " \cong " $\text{ht}^2 \geq C^2 \cdot \sigma$ " so $\mathbb{P}(\text{ht}(T) > \frac{k}{\sqrt{1-p_1}} \sqrt{\sigma}) \leq \exp(-\delta k^2)$

$$\text{Theorem: } \mathbb{P}(\text{ht}(T) > \frac{k}{\sqrt{1-p_1}} \sigma^{1/2}) \leq \exp(-\delta k^2)$$

Heuristic: GW trees satisfy $\text{wid}(T) \cdot \text{ht}(T) \approx \sigma$

Question: What is $\sup E[\text{wid}(T) \cdot \text{ht}(T) / \sigma]$?

Supremum over p where $T \sim \text{GW}(p)$.

Question: What is the behaviour of

$$\sup E[\text{wid}(T) \cdot \text{ht}(T) / \sigma \mid \sigma = n]$$

as a function of n ?

(At least $c \log n$; lower bound in article)

Galton-Watson Trees

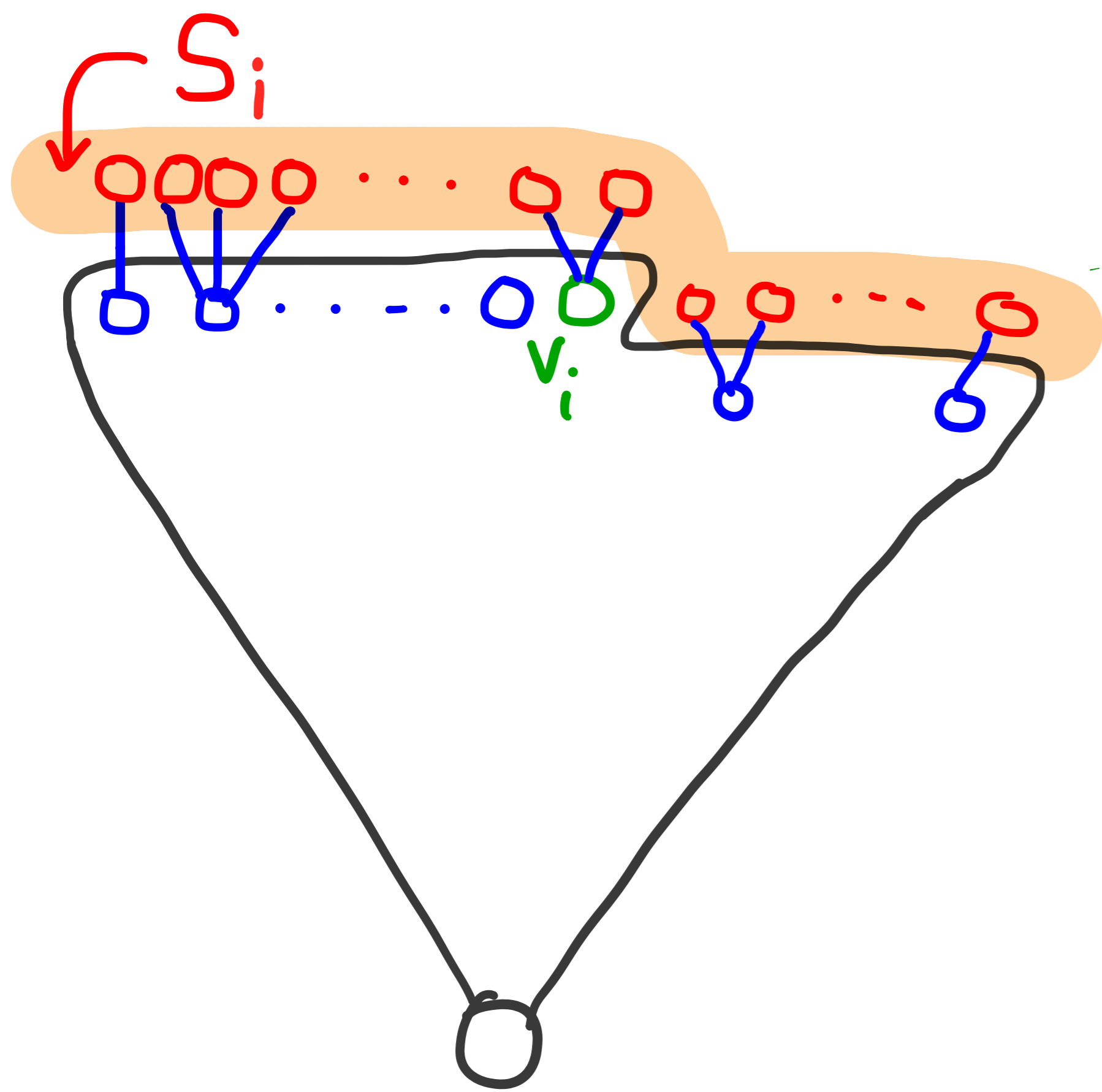
- Each node has random # of children
- Nodes reproduce independently

Construction

- $(C_i, i \geq 1)$ independent copies of a random variable C with $\sum_{k \geq 0} \mathbb{P}(C=k) = 1$.

Rule

The sequence $(C_i, i \geq 1)$ gives # children of nodes, in breadth-first search order



Halting Condition

For $i \geq 0$

nodes discovered by $\underbrace{\text{time } i}_{i \text{ steps of BFS}}$.

$$= 1 + \sum_{j=1}^i C_j$$

vertices explored by time $i = i$

$$\text{Let } S_i = 1 + \sum_{j=1}^i (C_j - 1)$$

= # nodes in "BFS queue" at time i
(discovered but not explored)

$S = (S_i, i \geq 0) = \text{BFS queue process.}$
(determines T)

Total # vertices = σ

$$= \inf \left\{ n : 1 + \sum_{j=1}^n C_j = n \right\} = \inf \left\{ n : S_n = 0 \right\}$$

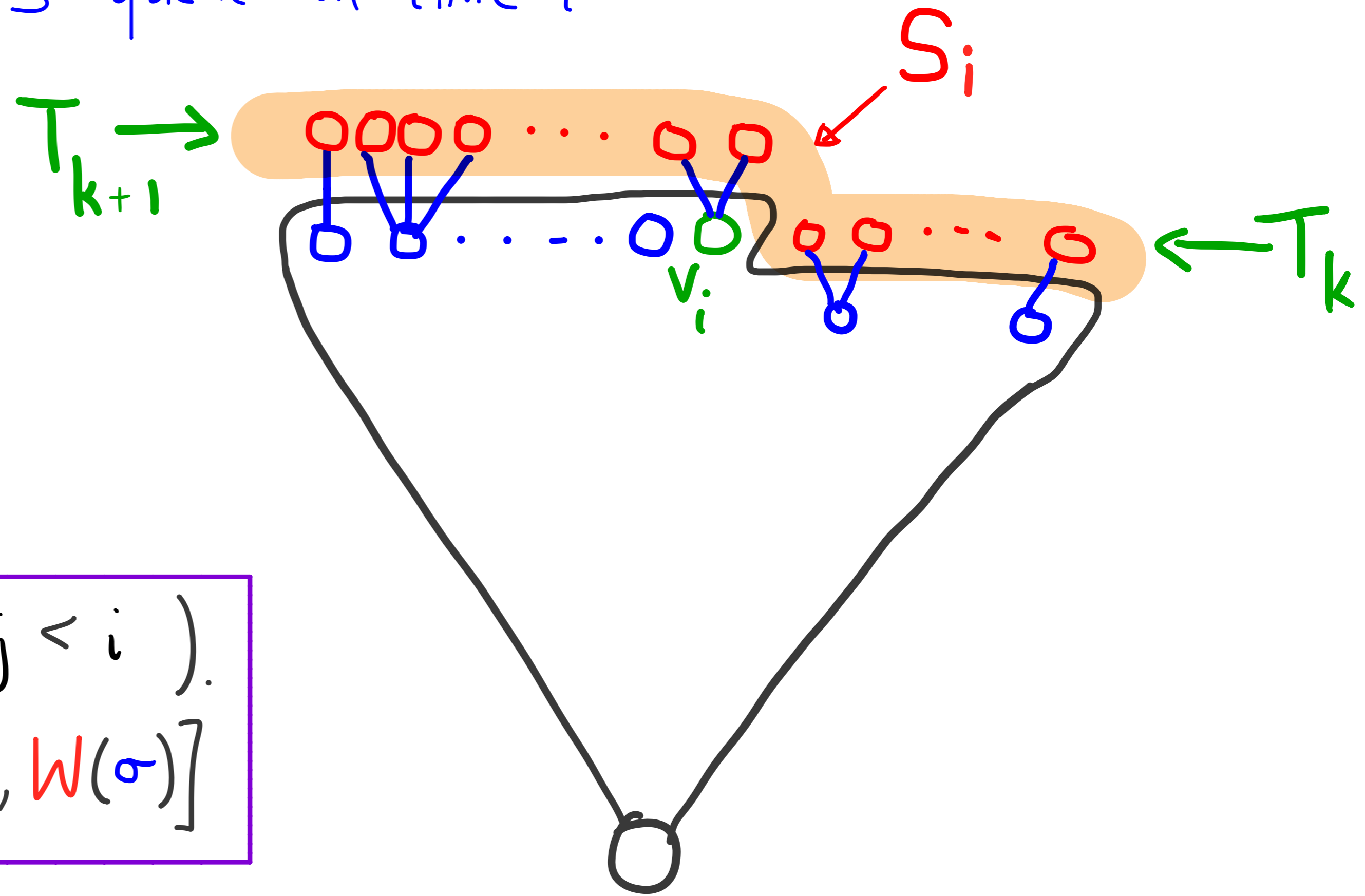
Setup

Let $S_i = 1 + \sum_{j=1}^i (C_j - 1) = \# \text{ nodes in "BFS queue" at time } i$

$$\mathbb{E} C \leq 1 \Rightarrow \mathbb{E}(S_{n+1} - S_n) = \mathbb{E} C - 1 \leq 0$$

$\sigma = \inf \{ n : S_n = 0 \}$
= first time no nodes left to explore

Prop: Let $W(i) = \max(S_j, 0 \leq j < i)$.
Then $\text{wid}(T) \in (W(\sigma)/2, W(\sigma)]$



Proof: During BFS on level k , "exploration queue" $\subset T_k \cup T_{k+1}$;
and $= T_k$ at start of level k . ■

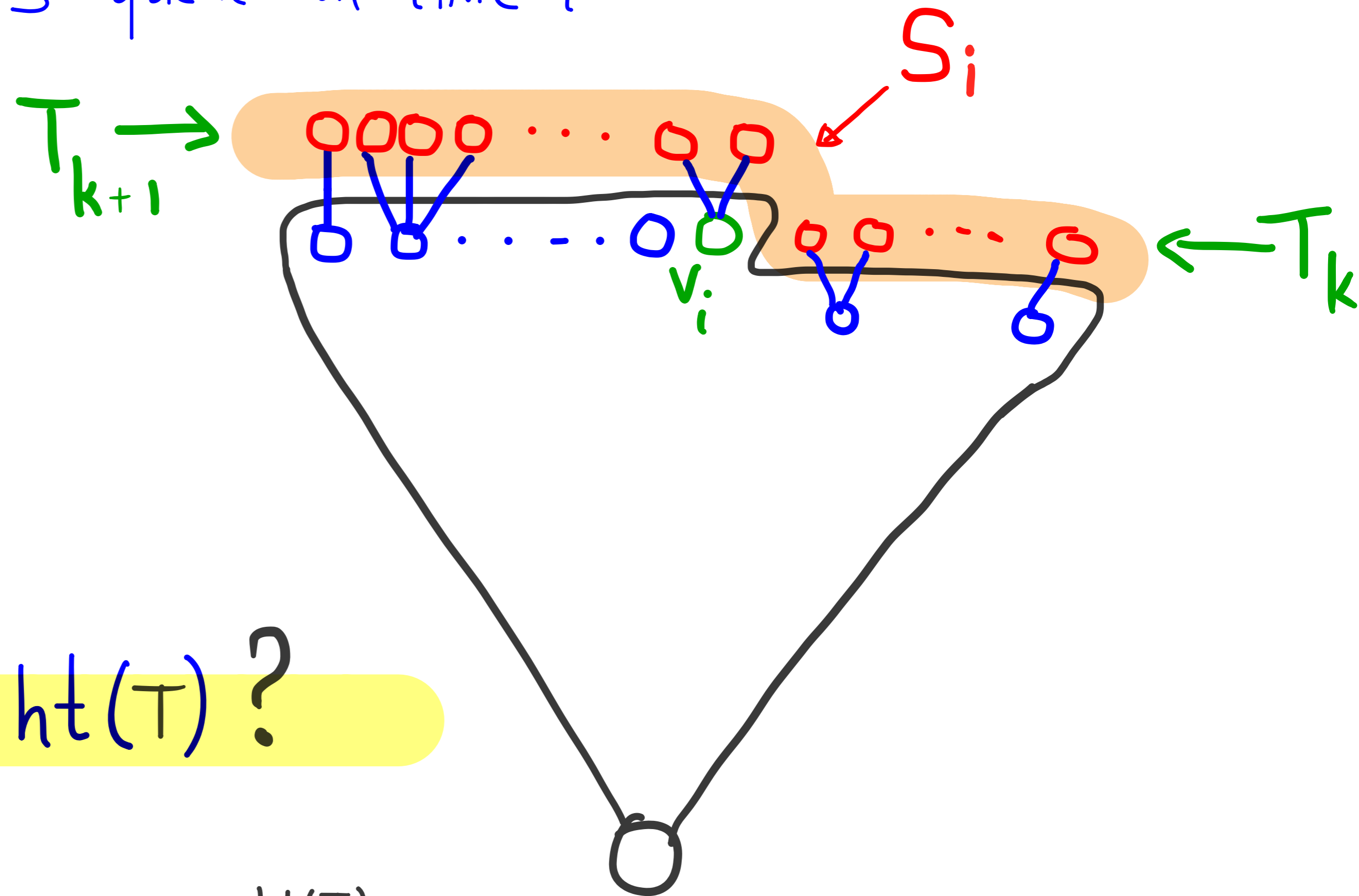
So $W(\sigma)$ is a good proxy for $\text{wid}(T)$.

Setup

Let $S_i = 1 + \sum_{j=1}^i (C_j - 1) = \# \text{ nodes in "BFS queue" at time } i$

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What can act as a proxy for $ht(T)$?

Idea: $ht(T) = \sum_{k=1}^{ht(T)} 1 = \sum_{k=1}^{ht(T)} \sum_{v \in T_k} \frac{1}{|T_k|} \approx \sum_{k=1}^{ht(T)} \sum_{v_i \in T_k} \frac{1}{S_i} = \sum_{i=1}^{\sigma} \frac{1}{S_i} =: H(\sigma)$

Prop:
 $ht(T) \leq 3H(\sigma)$

When $v_i \in T_k$ then $S_i \approx |T_k|$

Proof: If $v_i \in T_k$ then $S_i \leq |T_k \cup T_{k+1}|$, so $\sum_{v_i \in T_k} \frac{1}{S_i} \geq \frac{|T_k|}{|T_k \cup T_{k+1}|}$

If v_i is the j 'th node in T_k then $S_i \leq |T_k \cup T_{k+1}| - j$ so $\sum_{v_i \in T_k} \frac{1}{S_i} \geq \sum_{j \in |T_k|} \frac{1}{|T_k \cup T_{k+1}| - j} \geq \log \left(\frac{|T_k \cup T_{k+1}|}{|T_{k+1}|} \right)$

Combine these bounds, use that $S_0 = 1$, $S_\sigma = 0$ ■

Setup

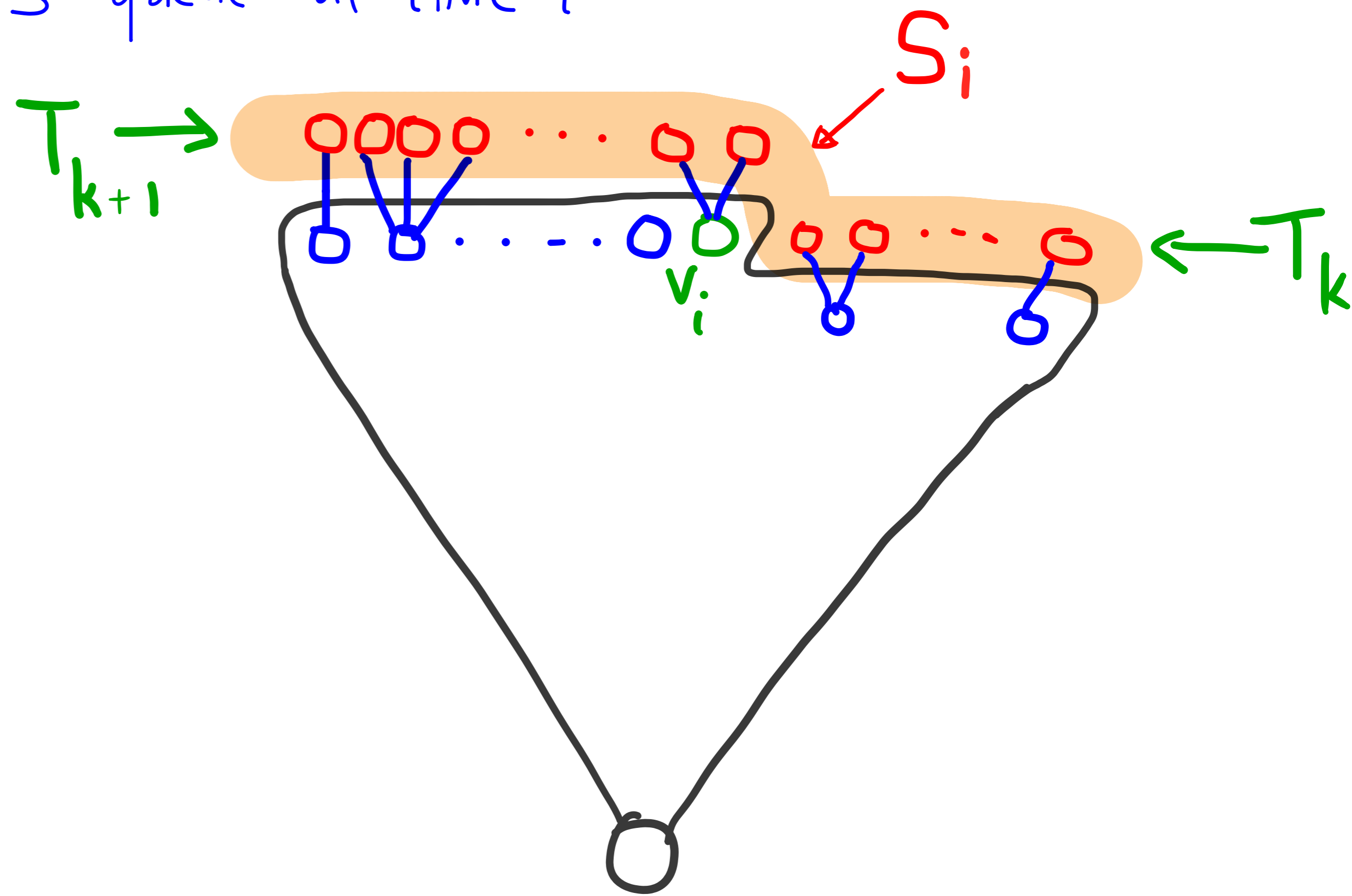
Let $S_i = 1 + \sum_{j=1}^i (C_j - 1) = \#$ nodes in "BFS queue" at time i

$$\mathbb{E} C \leq 1 \Rightarrow \mathbb{E}(S_{n+1} - S_n) = \mathbb{E} C - 1 \leq 0$$

$\sigma = \inf \{ n : S_n = 0 \}$
= first time no nodes left to explore

$$W(i) = \max(S_j, 0 \leq j < i).$$

$$H(i) = \sum_{j=1}^i \frac{1}{S_j}$$



Prop: $\text{wid}(T) \in (W(\sigma)/2, W(\sigma)]$

Prop:
 $\text{ht}(T) \leq 3H(\sigma).$

Corollary To prove $\mathbb{P}(\text{ht}(T) \geq \frac{k}{1-p_i} \cdot \text{wid}(T)) \leq \exp(-\delta k)$

suffices to prove $\mathbb{P}(H(\sigma) \geq \frac{k}{1-p_i} W(\sigma)) \leq e^{-\delta k}$.

$$W(i) = \max(S_j, 0 \leq j < i) \quad H(\sigma) = \sum_{j=1}^i \frac{1}{S_j} \quad \text{Aim: } \mathbb{P}(H(\sigma) \geq \frac{k}{1-p} W(\sigma)) \leq e^{-\delta k}$$

Key Tool: Decomposition into scales.

When $S_j \approx 2^l$ ("scale l "), have $H(j) - H(j-1) = \frac{1}{S_j} \approx \frac{1}{2^l}$

So bound (a) time to change scales,

(b) "# visits to scales" = $(V(l), l \geq 1)$

(a) Thm (Lévy; Doeblin; Kolmogorov; Rogozin; Le Cam; Esséen; Kesten):

With $p = \max p_i$, have

$$\max_k \mathbb{P}(S_n = k) \leq \frac{C p}{\sqrt{n(1-p)}} \quad C > 0 \text{ universal.}$$

"Any random walk spreads out over $\geq \sqrt{n}$ values by time n ."

Here $\sqrt{n} \approx 2^l$.

Thus if $S_i \in (2^l, 2^{l+1}]$, $\tau = \inf \{j > i : S_j \notin (2^{l-1}, 2^{l+2}]\}$

then $\frac{\tau-i}{4^l}$ has subexponential upper tail: $\mathbb{P}(\tau-i \geq k \cdot 4^l) \leq e^{-\delta k}$.

So $\frac{H(\tau) - H(i)}{2^l} \leq \frac{(\tau-i)/2^{l-1}}{2^l} = 2 \cdot \frac{\tau-i}{4^l}$ also has subexponential tail.

$$W(\sigma) = \max(S_i, 0 \leq i < \sigma) \quad H(\sigma) = \sum_{i=1}^{\sigma} \frac{1}{S_i} \quad \text{Aim: } \mathbb{P}(H(\sigma) \geq \frac{k}{1-p}, W(\sigma)) \leq e^{-\delta k}$$

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When $S_j \approx 2^l$ ("scale l "), have $H(j) - H(j-1) = \frac{1}{S_j} \approx \frac{1}{2^l}$

So bound (a) time to change scales,

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(b) $V(l) = \#$ visits to scale l

A visit to scale l starts at a time i with $S_i \in (2^l, 2^{l+1}]$;

ends at time $\tau = \inf \{j > i : S_j \notin (2^{l-1}, 2^{l+2}]\}$

$$W(\sigma) = \max(S_i, 0 \leq i < \sigma) \quad H(\sigma) = \sum_{i=1}^{\sigma} \frac{1}{S_i} \quad \text{Aim: } \mathbb{P}(H(\sigma) \geq \frac{k}{1-p}, W(\sigma)) \leq e^{-\delta k}$$

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So bound (a) time to change scales,

(b) "# visits to scales" = $(V(l), l \geq 1)$

(b) $V(l) = \# \text{ visits to scale } l = \#\{i : L_i = l\}$

Fact: Given that $V(l) \neq 0$, $V(l)$ dominated by sum of 2 $\text{Geom}(\frac{1}{2})$ r.v.s; $\Rightarrow \mathbb{P}(V(l) > k \mid V(l) > 0) \leq 2^{-k/2}$.

Proof: visits to scale l entail upcrossings of $[2^{l-1}, 2^{-l})$ or of $[2^{l+1}, 2^{l+2})$.

Both are hard since walk has non-positive drift. \blacksquare

$$W(\sigma) = \max(S_i, 0 \leq i < \sigma) \quad H(\sigma) = \sum_{i=1}^{\sigma} \frac{1}{S_i} \quad \text{Aim: } \mathbb{P}(H(\sigma) \geq \frac{k}{1-p} W(\sigma)) \leq e^{-\delta k}$$

Key Tool: Decomposition into scales.

When $S_j \approx 2^l$ ("scale l "), have $H(j) - H(j-1) = \frac{1}{S_j} \approx \frac{1}{2^l}$

So bound (a) time to change scales,

(b) "# visits to scales" = $(V(l), l \geq 1)$
 k 'th scale visited

(a) $\underbrace{\tau_{k+1} - \tau_k}_{\text{time to change scale}} \sim (2^{L_k})^2$ so $\mathbb{P}(H_{\tau_{k+1}} - H_{\tau_k} \geq \frac{x}{1-p} \cdot 2^l \mid L_k = l) \leq e^{-\delta x}$

(b) $V(l) = \# \text{ visits to scale } l$ $\mathbb{P}(V(l) > k \mid V(l) > 0) \leq 2^{-k/2}$

(a) + (b) \Rightarrow Total contribution of level l to height is

$$\begin{cases} 0 & \text{if } V(l) = 0. \\ O(2^l) & \text{with exp. tails if } V(l) > 0. \end{cases}$$

So $H(\sigma) \lesssim 2^{\max(l: V(l) \neq 0)} \lesssim W(\sigma)$

Remarks

- Stronger results if add info. about tails of degrees.

Ex • If $\mu(P) \leq 1$, $\text{Var}(P) = v \in (0, \infty)$ then $\exists x_0$ s.t. $\forall x \geq x_0$,

$$\mathbb{P}(\text{ht}(T) \geq Kx|T|^{1/2}) \leq \exp(-vx^2)$$

- If in fact $\mu(P) = 1$ then $\forall n \geq 1$,

$$\mathbb{P}(\text{ht}(T) \geq Kx|T|^{1/2} \mid |T| \geq n) \leq \exp(-vx^2)$$

- If $\exists \alpha \in (1, 2]$, $M > 0$ s.t. $\sum_{j>i} P_j > \frac{M}{i^\alpha} \quad \forall i \geq 1$, then $\forall x \geq 1$,

$$\mathbb{P}(\text{ht}(T) \geq \frac{Kx}{M(\alpha-1)} \cdot |T|^{\frac{\alpha-1}{2}}) \leq \exp(-x^\alpha)$$

- If $\mu(P) \leq 1$, $\text{var}(P) = \infty$ then $\forall \varepsilon > 0 \exists n_0$ s.t. $\forall x > 0, n \geq n_0$,

$$\mathbb{P}(\text{ht}(T) \geq x|T|^{1/2}, |T| \geq n) \leq \frac{x}{n^{1/2}} \exp(-x^2/\varepsilon)$$

Conjectures

- Conjecture: All this works even conditional on size of tree: $\mathbb{P}(\text{ht}(T) > A \cdot m \cdot \text{wid}(T) \mid \sigma = n) \leq \exp(-\delta m)$

NB: Here should have $\delta = \delta(p_0, p_1)$

- Conjecture: Binary trees are the tallest.

More specifically:

Consider random trees $T_{\vec{n}}$ with a fixed degree seq $\vec{n} = (n_i, i \geq 0)$

Here $n_i = \# \text{ nodes with } i \text{ children}$. With $\sum n_i = \sigma$, then $\sum i n_i = \sigma - 1$

Then to stochastically maximize $\text{ht}(T_{\vec{n}})$ among sequences with $n_0 = n, n_i = 0$, one should choose the seq. $(n, 0, n-1, 0, 0, 0, \dots)$

Part 2:

Binary trees stochastically maximize the height of a random node.

Conjecture: Binary trees are the tallest.

Consider random trees $T_{\vec{n}}$ with a fixed degree seq $\vec{n} = (n_i, i \geq 0)$.

Here $n_i = \#$ nodes of deg i .
 $\underbrace{\quad}_{\# \text{ children}}$

To stochastically maximize $ht(T_{\vec{n}})$ among sequences with $n_0 = n, n_i = 0$,

choose the seq. $\text{bin}(n) = (n, 0, n-1, 0, \dots)$

"Evidence."

Proposition: Let $\vec{n} = (n, 0, n_2, n_3, \dots)$, let $(T_{\vec{n}}, V)$ be a random marked tree with degree sequence \vec{n} .
 \uparrow distinguished node

Let $\text{bin}(n) = (n, 0, n-1, 0, 0, \dots)$, let $(T_{\text{bin}(n)}, W)$ be a random marked binary tree with n leaves.

Then $\text{height}(V) \preceq_{st} \text{height}(W)$.

Main idea of proof: grow the paths from the roots to V and to W one step at a time, compare conditional stopping probabilities.

(More details forthcoming.)

Warm-up:

binary trees, n leaves = $\frac{1}{2n-1} \binom{2n-1}{n}$

trees with degree seq. \vec{n} . With $|\vec{n}| := \sum n_i$ then get

$$\frac{1}{|\vec{n}|} \binom{|\vec{n}|}{n_i, i \geq 0} = \frac{1}{|\vec{n}|} \frac{|\vec{n}|!}{n_0! n_1! \dots}$$

$$= \frac{1}{|\vec{n}|} \cdot \# \text{ lattice walks, } n_i \text{ steps of size } i-1,$$

binary forests, n leaves, k connected components = $\frac{k}{2n-k} \binom{2n-k}{n}$

forests with degree seq. \vec{n} . With k trees, then get $\frac{k}{|\vec{n}|} \binom{|\vec{n}|}{n_i, i \geq 0}$

$$= \frac{k}{|\vec{n}|} \cdot \# \text{ lattice walks, } n_i \text{ steps of size } i-1,$$

marked forests with degree seq. \vec{n} , mark in last tree

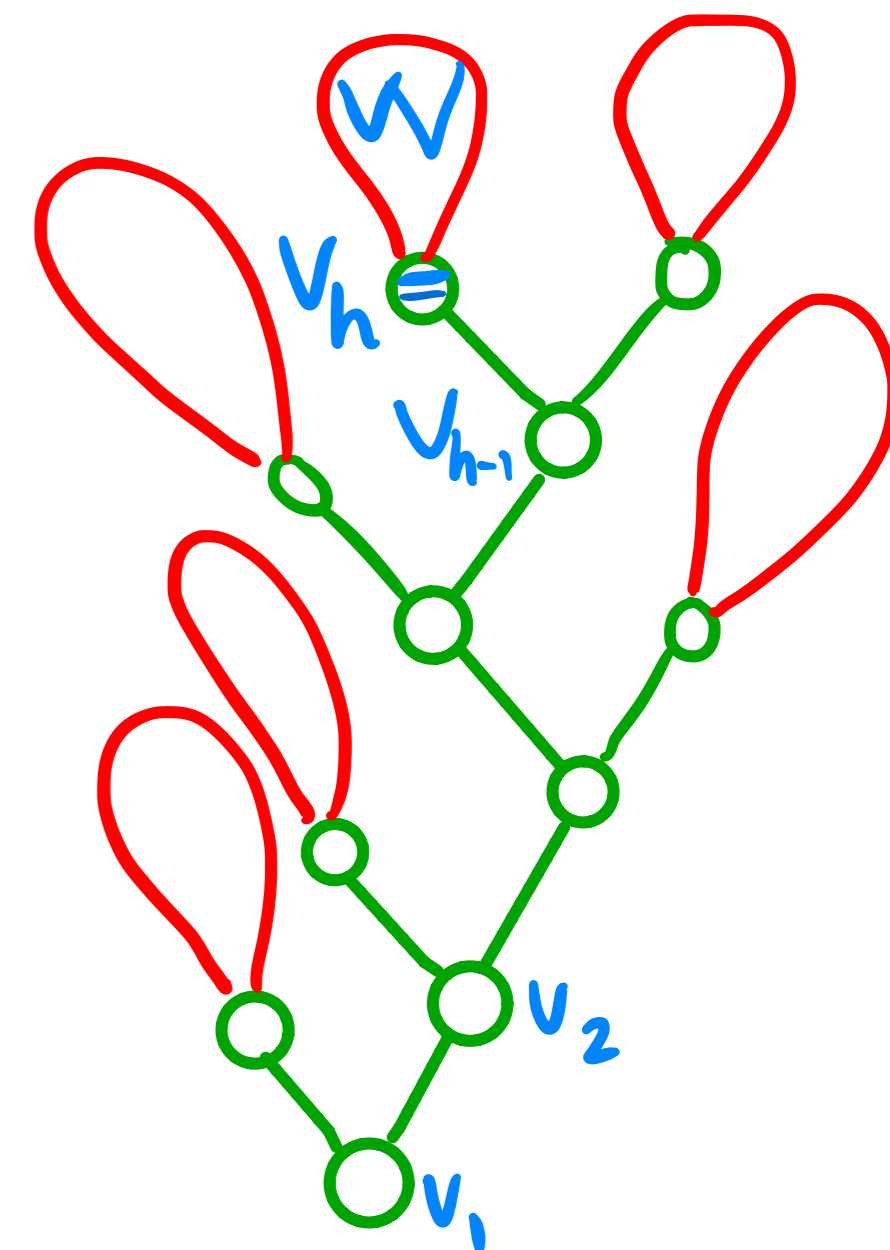
↑
distinguished
vertex

$$= |\vec{n}| \cdot \frac{1}{k} \cdot \# \text{ forests with degree seq. } \vec{n}$$

$$= \binom{|\vec{n}|}{n_i, i \geq 0} = \# \text{ lattice walks, } n_i \text{ steps of size } i-1$$

Trunks of trees

Let (T, W) be a random marked binary tree, n leaves.

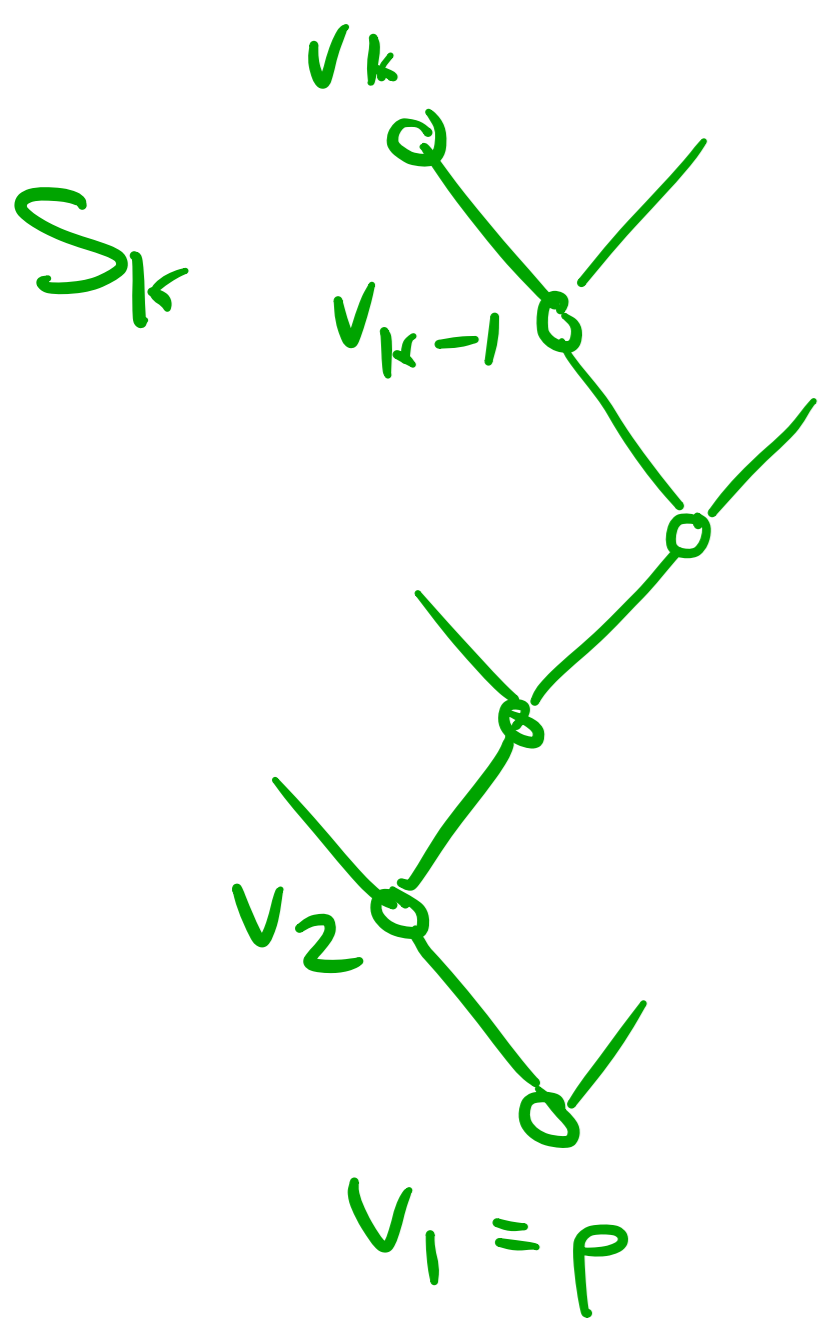


Trunk = path from root to marked vertex,

together with children of path vertices

Marked binary trees with n leaves, trunk containing S_k
(marked node in subtree rooted at v_k)

= # binary forests with n leaves, k trees, mark in last tree = $\binom{2n-k}{n}$



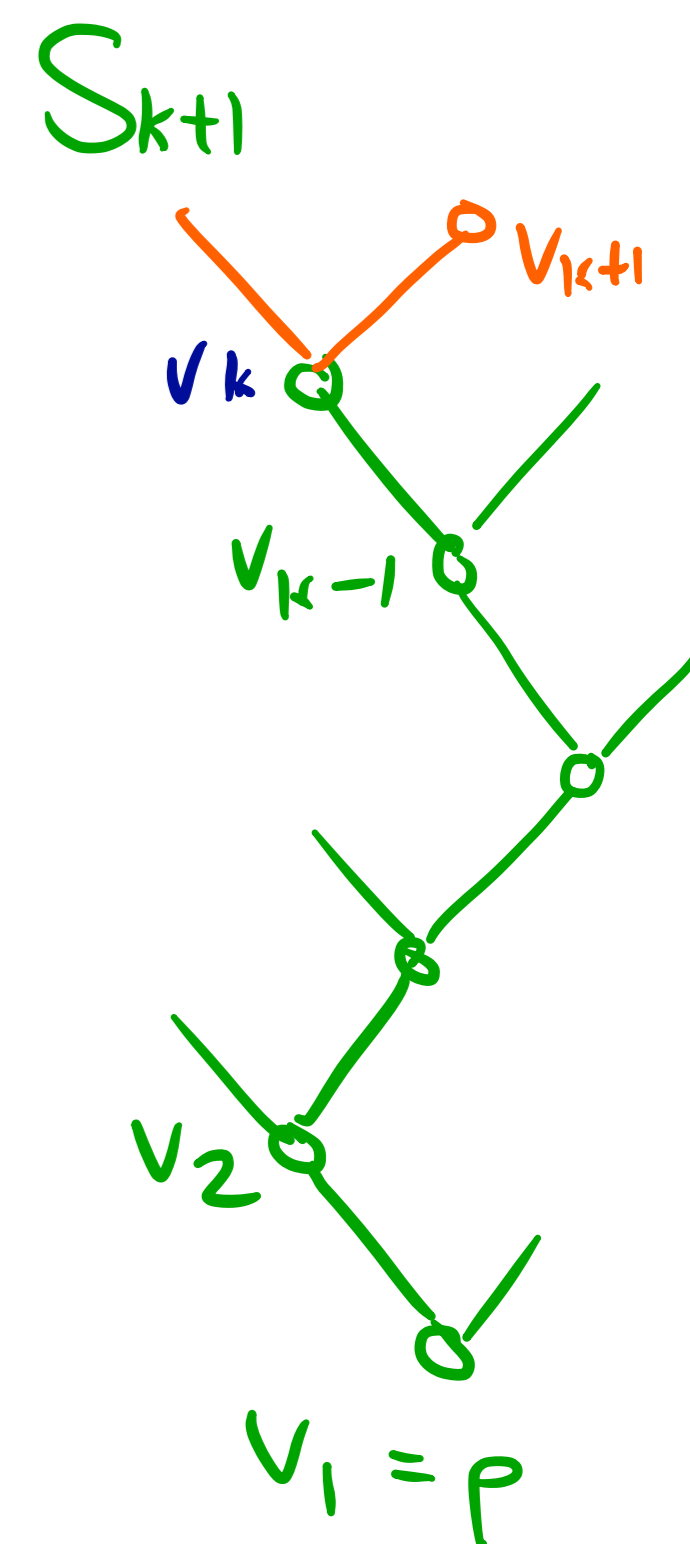
Marked binary trees with n leaves, trunk containing S_{k+1} , = $\binom{2n-k-1}{n}$

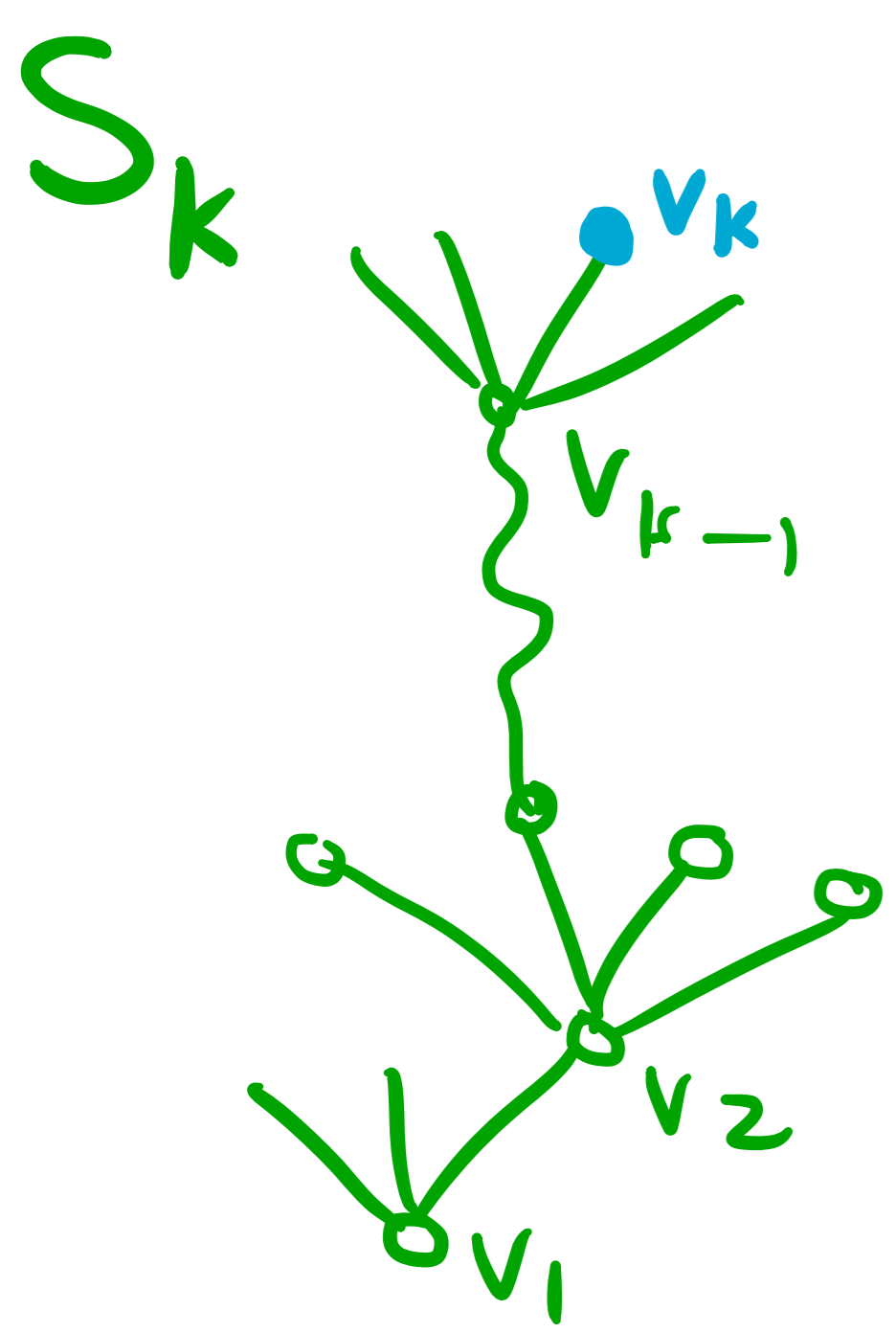
Ratio is $\frac{(2n-k-1)!}{n!(n-k-1)!} \cdot \frac{n!(n-k)!}{(2n-k)!} = \frac{n-k}{2n-k}$

Two possible choices for v_{k+1} (left or right)

So $P(W = v_k | \text{Trunk contains } S_k) = 1 - 2 \cdot \frac{n-k}{2n-k} = \frac{k}{2n-k}$

= $1/E(\text{Size of subtree above } v_k | T \text{ contains } S_k)$





Proposition

Let (T, V) be a random marked tree with degree sequence \vec{n} . Fix k and S_k with $\mathbb{P}(\text{Trunk contains } S_k) > 0$.

Then $\mathbb{P}(V = v_k \mid \text{Trunk contains } S_k)$

$$= 1 / \mathbb{E}(\text{Size of subtree above } v_k \mid T \text{ contains } S_k)$$

$$= \frac{1 + \sum_{i=1}^{k-1} (\deg(v_i) - 1)}{|\vec{n}| - (k-1)}$$

subtrees hanging from trunk.

vertices in these subtrees

NB: • With $\vec{n} = (n, 0, n_2, n_3, \dots)$ we have $|\vec{n}| \leq 2n - 1$

• Total # subtrees hanging from trunk is $1 + \sum_{i=1}^{k-1} (\deg(v_i) - 1) \geq k$,

so Proposition implies

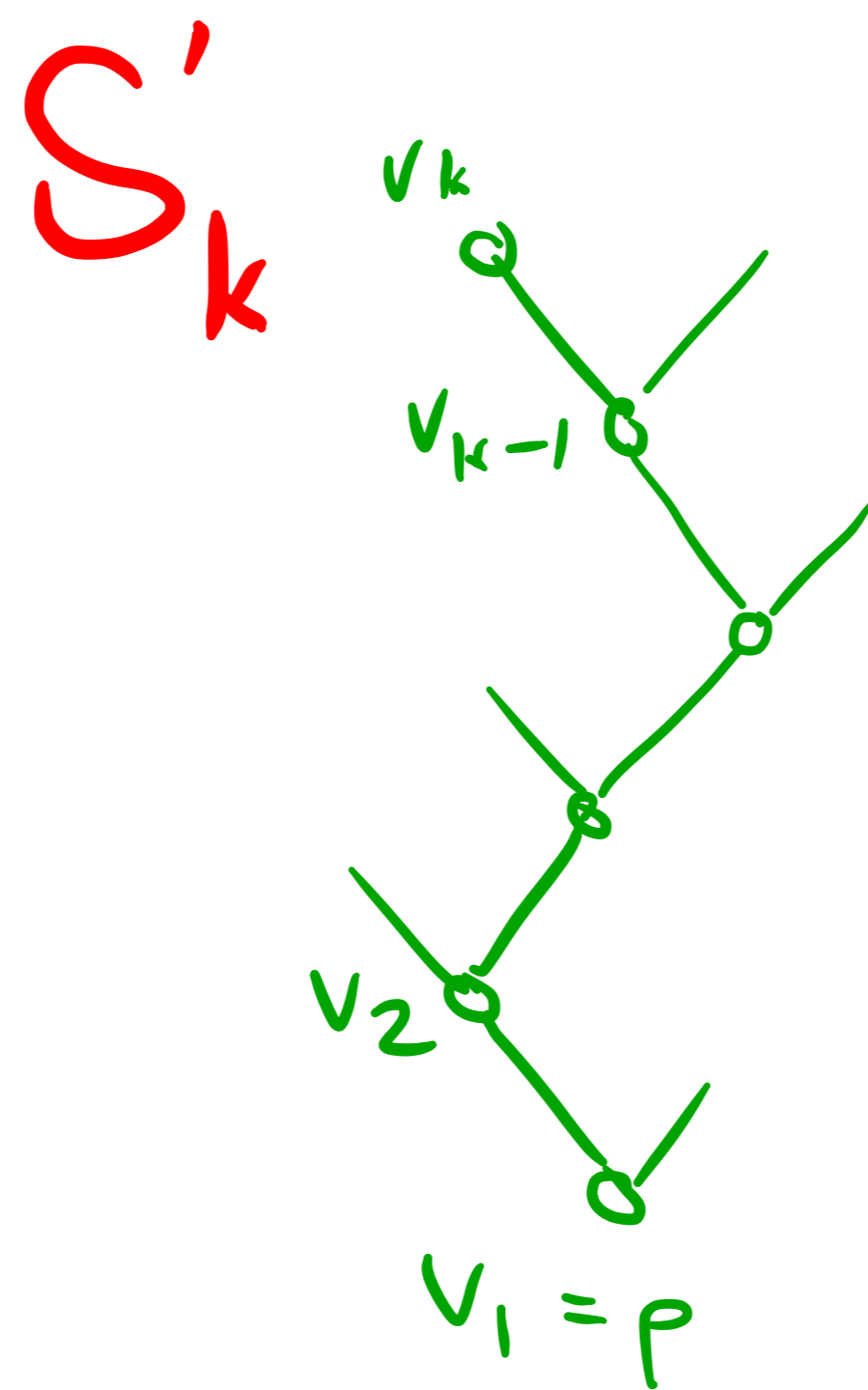
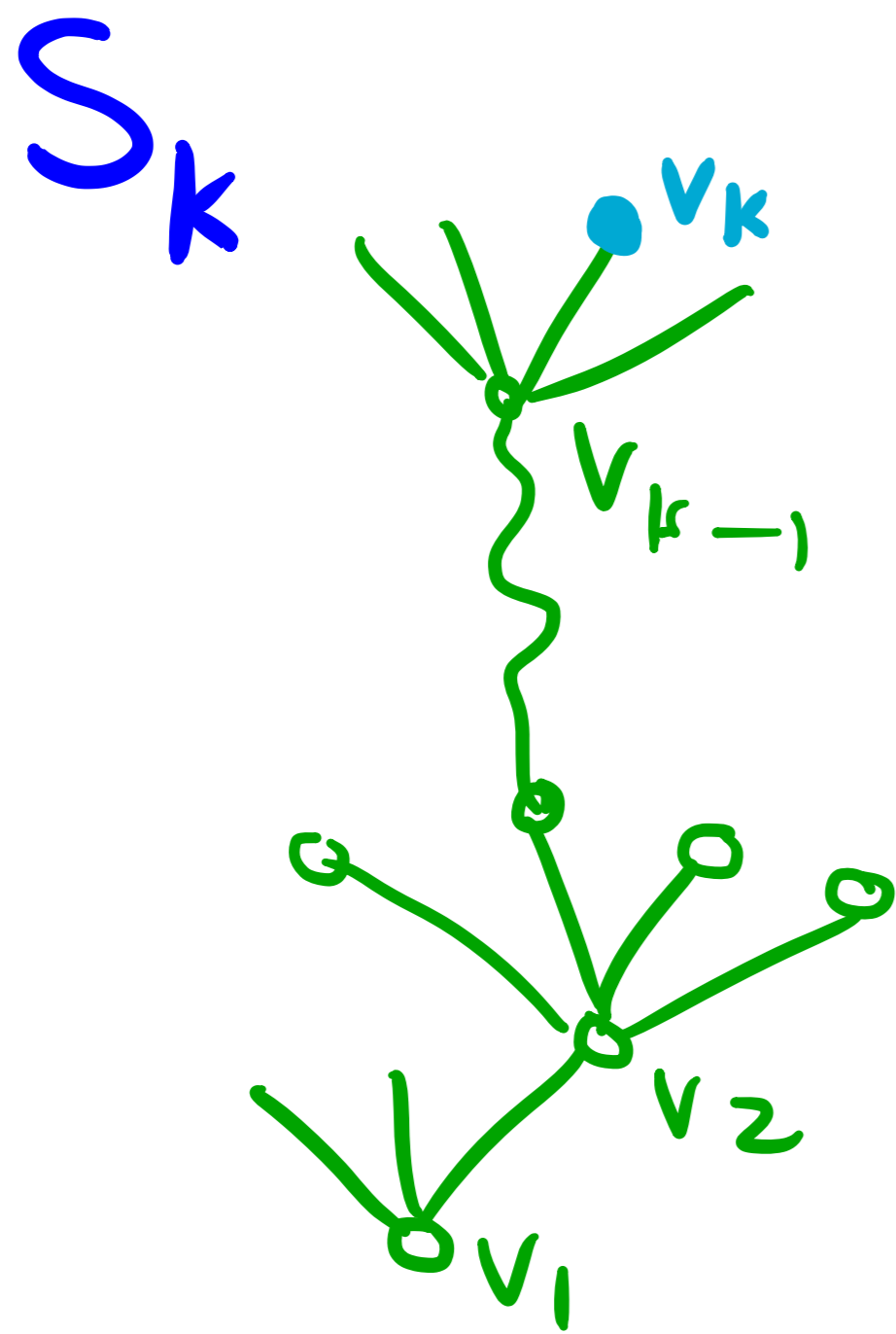
$$\mathbb{P}(V = v_k \mid T \text{ contains } S_k) \geq \frac{k}{2n - k}$$

Proposition: Let $\vec{n} = (n, 0, n_2, n_3, \dots)$, let $(T_{\vec{n}}, V)$ be a random marked tree with degree sequence \vec{n} .
↑ distinguished node

Let $\text{bin}(n) = (n, 0, n-1, 0, 0, \dots)$, let $(T_{\text{bin}(n)}, W)$ be a random marked binary tree with n leaves.

Then $\text{height}(V) \preceq_{\text{st}} \text{height}(W)$.

Proof: Fix S_k, S'_k possible "sub-trunks" for $T_{\vec{n}}, T_{\text{bin}(n)}$.



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Proof: Fix S_k, S'_k possible "sub-trunks" for $T_{\vec{n}}, T_{\text{bin}(n)}$.

In marked binary tree $\mathbb{P}(W = v_k \mid \text{Trunk contains } S'_k) = \frac{k}{2n-k}$

In marked tree with degree sequence \vec{n} , $\mathbb{P}(V = v_k \mid \text{Trunk contains } S_k) \geq \frac{k}{2n-k}$

So can couple step-by-step constructions of trunks so that at each step, the construction of marked tree with degree sequence \vec{n} is more likely to halt (if it has not already done so). \square

This sampling technique has other uses.

Theorem: There exists $c > 0$ such that the following holds. Fix any degree sequence $\vec{n} = (n_0, n_1, n_2, \dots)$, let $(T_{\vec{n}}, V)$ be a random marked tree with degree sequence \vec{n} . Write $\text{var}(\vec{n}) = \sum_{i>0} i(i-1) \frac{n_i}{|\vec{n}|}$. Then for all $x > 0$,

$$\mathbb{P}(\text{ht}(V) \geq x |\vec{n}|^{1/2}) \leq \exp(-c x^2 \text{var}(\vec{n})), \text{ and } \mathbb{E}[\text{ht}(V)] \leq \frac{1}{2c \text{var}(\vec{n})} \cdot |\vec{n}|^{1/2}.$$

Corollary: $\mathbb{E}[\text{wid}(T_{\vec{n}})] \geq \frac{c \text{var}(\vec{n})}{8} \cdot |\vec{n}|^{1/2}$

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Theorem: There exists $c > 0$ such that the following holds. Fix any degree sequence $\vec{n} = (n_0, n_1, n_2, \dots)$, let $(T_{\vec{n}}, V)$ be a random marked tree with degree sequence \vec{n} . Write $\text{var}(\vec{n}) = \sum_{i>0} i(i-1) \frac{n_i}{|\vec{n}|}$. Then for all $x > 0$,

$$\mathbb{P}(\text{ht}(V) \geq x |\vec{n}|^{1/2}) \leq \exp(-c x^2 \text{var}(\vec{n})), \text{ and } \mathbb{E}[\text{ht}(V)] \leq \frac{1}{2c \text{var}(\vec{n})} \cdot |\vec{n}|^{1/2}.$$

Corollary: $\mathbb{E}[\text{wid}(T_{\vec{n}})] \geq \frac{c \text{var}(\vec{n})}{8} \cdot |\vec{n}|^{1/2}$

Proof: By Markov's inequality,

$$\mathbb{P}\left[\text{ht}(V) \geq \frac{2}{c \text{var}(\vec{n})} \cdot |\vec{n}|^{1/2}\right] \leq \mathbb{P}\left[\text{ht}(V) \geq 4 \mathbb{E}[\text{ht}(V)]\right] \leq \frac{1}{4}$$

$$\text{Also } \mathbb{P}\left[\text{ht}(V) \geq \frac{2}{c \text{var}(\vec{n})} \cdot |\vec{n}|^{1/2}\right] \geq \frac{1}{2} \cdot \mathbb{P}\left[\#\{u \in T_{\vec{n}} : \text{ht}(u) \geq \frac{2}{c \text{var}(\vec{n})} \cdot |\vec{n}|^{1/2}\} \geq \frac{|\vec{n}|}{2}\right]$$

$$\text{So } \mathbb{P}\left[\#\{u \in T_{\vec{n}} : \text{ht}(u) \geq \frac{2}{c \text{var}(\vec{n})} \cdot |\vec{n}|^{1/2}\} \geq \frac{|\vec{n}|}{2}\right] \leq \frac{1}{2}$$

This sampling technique has other uses.

Theorem: There exists $c > 0$ such that the following holds. Fix any degree sequence $\vec{n} = (n_0, n_1, n_2, \dots)$, let $(T_{\vec{n}}, V)$ be a random marked tree with degree sequence \vec{n} . Write $\text{var}(\vec{n}) = \sum_{i>0} i(i-1) \frac{n_i}{|\vec{n}|}$. Then for all $x > 0$,

$$\mathbb{P}(\text{ht}(V) \geq x |\vec{n}|^{1/2}) \leq \exp(-c x^2 \text{var}(\vec{n})), \text{ and } \mathbb{E}[\text{ht}(V)] \leq \frac{1}{2c \text{var}(\vec{n})} \cdot |\vec{n}|^{1/2}.$$

Corollary: $\mathbb{E}[\text{wid}(T_{\vec{n}})] \geq \frac{c \text{var}(\vec{n})}{8} \cdot |\vec{n}|^{1/2}$

Proof:
(continued) $\mathbb{P}\left[\#\left\{u \in T_{\vec{n}} : \text{ht}(u) \geq \frac{2}{c \text{var}(\vec{n})} \cdot |\vec{n}|^{1/2}\right\} \geq \frac{|\vec{n}|}{2}\right] \leq \frac{1}{2}$

But if $\exists \geq \frac{|\vec{n}|}{2}$ nodes at height $\leq \frac{2}{c \text{var}(\vec{n})} \cdot |\vec{n}|^{1/2}$

then $\text{wid}(T_{\vec{n}}) \geq \frac{c \text{var}(\vec{n})}{4} \cdot |\vec{n}|^{1/2}$ by the pigeonhole principle.

So $\mathbb{E}[\text{wid}(T_{\vec{n}})] \geq \frac{c \text{var}(\vec{n})}{4} \cdot |\vec{n}|^{1/2} \cdot \frac{1}{2}$ ■

Results

2,4 Tight;
1,3,5 off
by $(\log n)^{1/2}$
factor.

Theorem 1: If $\mathbb{E} X = 1$, $\mathbb{E} X^2 = \infty$ then $ht(\mathcal{T}_\omega^{(n)}) / (n \log n)^{1/2} \xrightarrow{P} 0$

Theorem 2: If $\mathbb{E} X = 1$, $\mathbb{E} X^2 = \infty$ then $wid(\mathcal{T}_\omega^{(n)}) / n^{1/2} \xrightarrow{P} \infty$

Theorem 3: If $\mathbb{E} X < 1$ then $ht(\mathcal{T}_\omega^{(n)}) / (n \log n)^{1/2} \xrightarrow{P} 0$

Theorem 4: If $\mathbb{E} X < 1$ then $wid(\mathcal{T}_\omega^{(n)}) / n^{1/2} \xrightarrow{P} \infty$

Theorem 5: $ht(\mathcal{T}_\omega^{(n)}) / (n \log n)^{1/2}$ has uniform subgaussian tails

Conjectures of
Svante Janson
from

"Simply-generated trees, conditioned Galton-Watson trees,
random allocations and condensation",

Prob. Surveys Vol 9 (2012), pages 103-252

Conjecture 21.5: If $\mathbb{E} X = 1$, $\mathbb{E} X^2 = \infty$ then $ht(\mathcal{T}_\omega^{(n)}) / n^{1/2} \xrightarrow{P} 0$

Conjecture 21.6: If $\mathbb{E} X = 1$, $\mathbb{E} X^2 = \infty$ then $wid(\mathcal{T}_\omega^{(n)}) / n^{1/2} \xrightarrow{P} \infty$

Problem 21.7: Does $\mathbb{E} X < 1$ imply that $ht(\mathcal{T}_\omega^{(n)}) / n^{1/2} \xrightarrow{P} 0$?

Problem 21.8: Does $\mathbb{E} X < 1$ imply that $wid(\mathcal{T}_\omega^{(n)}) / n^{1/2} \xrightarrow{P} \infty$?

Problem 21.9: Does $ht(\mathcal{T}_\omega^{(n)}) / n^{1/2}$ have uniform subgaussian tails?

Can pass results to simply generated trees by "averaging over the degree sequence".

Svante
Janson:
Simply
Generated
Trees,
conditioned
Galton-
Watson
Trees,
Random
Allocations
and
Condens-
ation.

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Conjecture 21.5. If $\nu = 1$ and $\sigma^2 = \infty$, then $H(\mathcal{T}_n)/\sqrt{n} \xrightarrow{P} 0$. $\sqrt{n \log n}$ ✓

Conjecture 21.6. If $\nu = 1$ and $\sigma^2 = \infty$, then $W(\mathcal{T}_n)/\sqrt{n} \xrightarrow{P} \infty$. ✓

Problem 21.7. Does $\nu < 1$ imply that $H(\mathcal{T}_n)/\sqrt{n} \xrightarrow{P} 0$? $\sqrt{n \log n}$ ✓

Problem 21.8. Does $\nu < 1$ imply that $W(\mathcal{T}_n)/\sqrt{n} \xrightarrow{P} \infty$? ✓

Furthermore, still in the case $\nu \geq 1$, $\sigma^2 < \infty$, Addario-Berry, Devroye and Janson [1] have shown sub-Gaussian tail estimates for the height and width

$$\mathbb{P}(H(\mathcal{T}_n) \geq x\sqrt{n}) \leq Ce^{-cx^2}, \quad (21.12)$$

$$\mathbb{P}(W(\mathcal{T}_n) \geq x\sqrt{n}) \leq Ce^{-cx^2}, \quad (21.13)$$

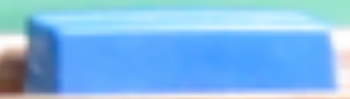
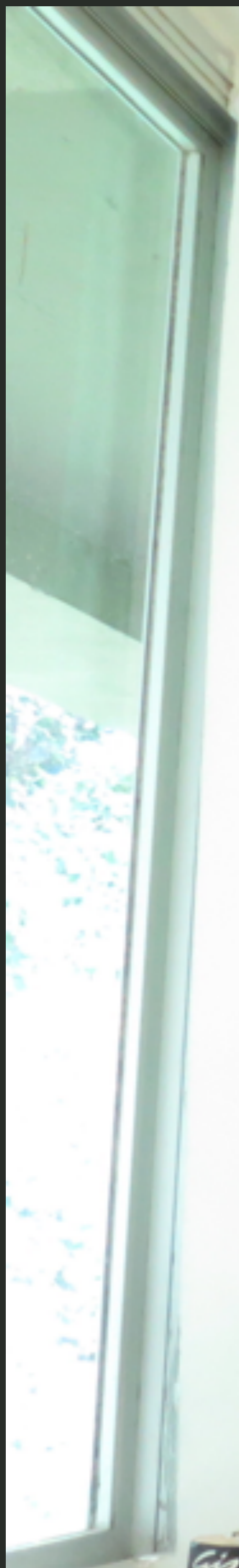
uniformly in all $x \geq 0$ and $n \geq 1$ (with some positive constants C and c depending on π and thus on \mathbf{w}). In view of (21.11), we cannot expect (21.13) to hold when $\sigma^2 = \infty$ (or when $\nu < 1$), but we see no reason why (21.12) cannot hold; (21.10) suggests that $H(\mathcal{T}_n)$ typically is smaller when $\sigma^2 = \infty$.

Problem 21.9. Does (21.12) hold for any weight sequence \mathbf{w} (with C and c depending on \mathbf{w} , but not on x or n)? Up to $\sqrt{n \log n}$ factor ✓

IMPACASTORINA - PRINCEPES

IMPA-NWL - NET

Thank you!



- Claimed theorem with dependence only on p_i proved it with dependence on $p = \max p_i$.

Fix: requires more careful "dispersion" bound for our setting.

(Idea: If subcritical then $p_{\max} = p_0$ or p_i ; if p_0 close to 1 then either very subcritical or make large jumps.)